

A Class of Fractional Stochastic Differential Equations

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Abstract. In this paper we consider the fractional case of a class of stochastic differential equations that has many important applications. Based on an approximation approach we solve the equation with polynomial drift and fractional noise. An explicit solution is found and some applications are given.

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1. Introduction

The fractional Brownian motion (fBm) of Hurst parameter $H \in (0, 1)$ is a centered Gaussian process $B^H = \{B_t^H, t \geq 0\}$ with the covariance function $R_H(t, s) = E[B_t^H B_s^H]$

$$R_H(t, s) = \frac{1}{2} (t^{2H} + s^{2H} - |t - s|^{2H}).$$

In the case where $H = \frac{1}{2}$, the process B^H is a standard Brownian motion. If $H \neq \frac{1}{2}$, B^H is neither a semimartingale nor a Markov process and the stochastic calculus developed by Itô cannot be applied. There are various approaches to fractional stochastic calculus by using some difficult tools such as: Malliavin calculus (see, for instance [1, 4]), theory of Wick product [5]. However, it is not easy to find explicit solutions from these methods for many practical problems.

In this paper using an approximation approach in $L^2(\Omega)$, we investigate a class of fractional stochastic equations of the form

$$dX_t = (aX_t^n + bX_t)dt + cX_t dB_t, \quad (1.1)$$

where $n \in \mathbb{N}, n \geq 2$ and B_t is a fractional Brownian motion of Liouville form that is defined below.

This equation is a generalization of many important equations such as: the Black-Scholes equation in mathematical finance, the Ginzburg-Landau equation in theoretical physics, the Verlhust equation in population study.

In [8] Mandelbrot has given a representation of B^H of the form:

$$B_t^H = \frac{1}{\Gamma(1+\alpha)} [U_t + B_t],$$

where $U_t = \int_{-\infty}^0 ((t-s)^\alpha - (-s)^\alpha) dW_s$, $B_t = \int_0^t (t-s)^\alpha dW_s$ and $\alpha = H - \frac{1}{2}$. B_t is a process possessing main properties of B_t^H such as of long memory and it is called a fractional Brownian motion of Liouville form [2, 6].

It is known that B_t is approximated in $L^2(\Omega)$ by stochastic processes

$$\tilde{B}_t = \int_0^t (t-s+\varepsilon)^\alpha dW_s,$$

where \tilde{B}_t is a semimartingale and the convergence is uniform in $t \in [0, T]$.

The paper is organized as follows: In Sec. 2 we recall some important results from the approximation approach and formulate our approximation problem. Section 3 contains main results of this paper. In Sec. 4 some applications to finance and physics are introduced.

2. An Approximation Method

Our method is based on a result on approximation of the fractional process $B_t = \int_0^t (t-s)^\alpha dW_s$ by semimartingales given in [9] that we recall below:

For every $\varepsilon > 0$, as in [1] we define:

$$\tilde{B}_t = \int_0^t (t-s+\varepsilon)^\alpha dW_s, \quad \alpha = H - \frac{1}{2} \in \left(-\frac{1}{2}, \frac{1}{2}\right). \quad (2.1)$$

Then we have

Theorem 2.1. I. The process $\{\tilde{B}_t, t \geq 0\}$ is a semimartingale. Moreover

$$\tilde{B}_t = \alpha I(t) + \varepsilon^\alpha W_t, \tag{2.2}$$

where $\varphi_\varepsilon(t) = \int_0^t (t - s + \varepsilon)^{\alpha-1} dW_s$ and $I(t) = \int_0^t \varphi_\varepsilon(s) ds$.

II. The process \tilde{B}_t converges to B_t in $L^2(\Omega)$ when ε tends 0. This convergence is uniform with respect to $t \in [0, T]$, i.e.:

$$\lim_{\varepsilon \rightarrow 0} \sup_{0 \leq t \leq T} \|\tilde{B}_t - B_t\|_2 = 0. \tag{2.3}$$

Proof. Refer to [9]. ■

Corollary 2.2. For any $p \geq 1$, the process \tilde{B}_t uniformly converges in t to B_t in $L^p(\Omega)$ when ε tends 0.

Proof. Noting that $\tilde{B}_t \sim \mathcal{N}(0, \tilde{\sigma}_t^2)$ and $B_t \sim \mathcal{N}(0, \sigma_t^2)$

$$\tilde{\sigma}_t^2 := E|\tilde{B}_t|^2 = \frac{(t + \varepsilon)^{2H} - \varepsilon^{2H}}{2H} \quad \text{and} \quad \sigma_t^2 := E|B(t)|^2 = \frac{t^{2H}}{2H}.$$

Since $t \in [0, T]$ and $\varepsilon \rightarrow 0^+$, it follows that both $\tilde{\sigma}_t^2$ and σ_t^2 are bounded by some positive constant. Thus, the Gaussian processes $B_\varepsilon(t)$ and $B(t)$ have null means and finite variance. Hence they have finite moments of any order.

If $1 \leq p \leq 2$ then by applying the Lyapunov inequality, we obtain

$$(E|\tilde{B}_t - B_t|^p)^{\frac{1}{p}} \leq (E|\tilde{B}_t - B_t|^2)^{\frac{1}{2}} \rightarrow 0 \quad \text{when } \varepsilon \rightarrow 0.$$

If $p > 2$ then it follows from Hölder inequality

$$E|X + Y|^p \leq 2^{p-1}(E|Y|^p + E|Y|^p) \quad \text{for any } p \geq 1$$

that

$$\begin{aligned} E|\tilde{B}_t - B_t|^p &\leq (E|\tilde{B}_t - B_t|^2)^{\frac{1}{2}} (E|\tilde{B}_t - B_t|^{2p-2})^{\frac{1}{2}} \\ &\leq \|\tilde{B}_t - B_t\|_2 \left[2^{2p-3} (E|\tilde{B}_t|^{2p-2} + E|B_t|^{2p-2}) \right]^{\frac{1}{2}}. \end{aligned}$$

Since \tilde{B}_t and B_t have moments of any order so there exists some constant M_p depending only p such that

$$E|\tilde{B}_t - B_t|^p \leq M_p \|\tilde{B}_t - B_t\|_2.$$

The proof is thus complete. ■

Next, let us consider the following fractional differential equation in a complete probability space (Ω, \mathcal{F}, P)

$$\begin{cases} dX_t = (a X_t^n + b X_t) dt + c X_t dB_t \\ X_t|_{t=0} = X_0 \end{cases} \tag{2.4}$$

or

$$\begin{cases} X_t - X_0 = \int_0^t (a X_s^n + b X_s) ds + c \int_0^t X_s dB_s \\ X_t|_{t=0} = X_0, \end{cases} \tag{2.5}$$

where the stochastic integral $\int_0^t X_s dB_s$ will be defined as the L^2 -limit of $\int_0^t X_s d\tilde{B}_s$ when $\varepsilon \rightarrow 0$, if it exists. The initial value X_0 is a measurable random variable independent of $\{B_t : 0 \leq t \leq T\}$.

As we said in Introduction, for the fractional stochastic calculus it is not good enough to find explicit solutions of fractional stochastic differential equations. In order to avoid this difficulty and moreover, because $\tilde{B}_t \rightrightarrows B_t$ it will be fully natural to approximate (2.4) by the following equation

$$\begin{cases} d\tilde{X}_t = (a \tilde{X}_t^n + b \tilde{X}_t) dt + c \tilde{X}_t d\tilde{B}_t \\ \tilde{X}_t|_{t=0} = X_0. \end{cases} \tag{2.6}$$

And then the solution of equation (2.4) will be the limit in $L^2(\Omega)$ of the solution of (2.6) when $\varepsilon \rightarrow 0$.

3. Main Results

The equation (2.6) is a stochastic differential equation driven by a semimartingale with a polynomial drift and a constant volatility. So the existence and uniqueness of its solution are assured. Using formula (2.2) we can rewrite equation (2.6) as follows

$$\begin{cases} d\tilde{X}_t = (a \tilde{X}_t^n + b \tilde{X}_t + c \alpha \varphi^\varepsilon(t) \tilde{X}_t) dt + c \varepsilon^\alpha \tilde{X}_t dW_t \\ \tilde{X}_t|_{t=0} = X_0. \end{cases} \tag{3.1}$$

We have the following theorem.

Theorem 3.1 *The solution of equation (2.6) can be explicitly given by*

$$\tilde{X}_t = e^{(b - \frac{1}{2}c^2\varepsilon^{2\alpha})t + c\tilde{B}_t} \left(X_0^{1-n} + (1-n)a \int_0^t e^{(n-1)((b - \frac{1}{2}c^2\varepsilon^{2\alpha})s + c\tilde{B}_s)} ds \right)^{\frac{1}{1-n}}.$$

Proof. Put $Y_t = e^{-c\varepsilon^\alpha W_t}$. According to the Itô formula we have

$$dY_t = Y_t \left(\frac{1}{2} c^2 \varepsilon^{2\alpha} dt - c \varepsilon^\alpha dW_t \right). \tag{3.2}$$

We consider $Z_t = \tilde{X}_t Y_t$ and we get by applying the integration of parts formula

$$\begin{aligned} dZ_t &= \tilde{X}_t dY_t + Y_t d\tilde{X}_t - c^2 \varepsilon^{2\alpha} \tilde{X}_t Y_t dt \\ &= \left\{ a e^{(n-1)c\varepsilon^\alpha W_t} (Z_t)^n + (b + c\alpha \varphi^\varepsilon(t) - \frac{1}{2} c^2 \varepsilon^{2\alpha}) Z_t \right\} dt. \end{aligned} \tag{3.3}$$

This is an ordinary Bernoulli equation of the form

$$Z'_t = P(t)Z_t^n + Q(t)Z_t$$

and the solution Z_t is given by

$$Z_t = e^{\int_0^t Q(u)du} \left(Z_0 + \int_0^t (1-n)P(s)e^{-(n-1)\int_0^s Q(u)du} ds \right)^{\frac{1}{1-n}},$$

where $P(t) = a e^{(n-1)c\varepsilon^\alpha W_t}$, $Q(t) = b - \frac{1}{2} c^2 \varepsilon^{2\alpha} + c\alpha \varphi^\varepsilon(t)$ and $\int_0^t Q(u)du = (b - \frac{1}{2} c^2 \varepsilon^{2\alpha})t + c\alpha I(t)$.

Hence the solution Z_t of equation (3.3) can be expressed as

$$Z_t = e^{(b-\frac{1}{2}c^2\varepsilon^{2\alpha})t+c\alpha I(t)} \left(Z_0 + (1-n)a \int_0^t e^{(n-1)((b-\frac{1}{2}c^2\varepsilon^{2\alpha})s+c\alpha I(s)+c\varepsilon^\alpha W(s))} ds \right)^{\frac{1}{1-n}}.$$

Combining the last expression and $\tilde{B}_t = \alpha I(t) + \varepsilon^\alpha W_t$ we obtain a solution of the approximation equation (2.6)

$$\tilde{X}_t = e^{(b-\frac{1}{2}c^2\varepsilon^{2\alpha})t+c\tilde{B}_t} \left(X_0^{1-n} + (1-n)a \int_0^t e^{(n-1)((b-\frac{1}{2}c^2\varepsilon^{2\alpha})s+c\tilde{B}_s)} ds \right)^{\frac{1}{1-n}}.$$

The proof is thus complete. ■

Theorem 3.2. *Suppose that X_0 is a random variable such that $X_0 > 0$ a.s. and $E[X_0^{2n}] < \infty$. If $H > \frac{1}{2}$ and $a \leq 0$ then the stochastic process X_t^* defined by*

$$X_t^* = e^{bt+cB_t} \left(X_0^{1-n} + (1-n)a \int_0^t e^{(n-1)(bs+cB_s)} ds \right)^{\frac{1}{1-n}} \tag{3.4}$$

is the limit in $L^2(\Omega)$ of \tilde{X}_t . This limit is uniform with respect to $t \in [0, T]$.

Proof. Put $\theta_\varepsilon(t) = e^{(b-\frac{1}{2}c^2\varepsilon^{2\alpha})t+c\tilde{B}_t}$ and $\theta(t) = e^{bt+cB_t}$. Then it is clear that for each $m \geq 1$ there exists a finite constant $M_m > 0$ such that $E[\theta_\varepsilon^m(t)] \leq M_m$, $E[\theta^m(t)] \leq M_m$ for every $t \in [0, T]$. Indeed,

$$E[\theta^m(t)] = e^{mbt} E[e^{m c B_t}] = e^{mbt} e^{\frac{1}{2}(m c b_t)^2} = e^{mbt+\frac{1}{2}m^2c^2b_t^2} < \infty,$$

and

$$E[\theta_\varepsilon^m(t)] = e^{m(b-\frac{1}{2}c^2\varepsilon^{2\alpha})t+\frac{1}{2}m^2c^2\tilde{b}_t^2} < \infty.$$

Moreover, by applying Hölder inequality we have the following estimates for any $m, k \geq 1$:

$$E[\theta_\varepsilon^m(t) \theta^k(t)] \leq (E|\theta_\varepsilon(t)|^{2m})^{\frac{1}{2}} (E|\theta(t)|^{2k})^{\frac{1}{2}}.$$

So there exists a finite constant $M_{m,k} > 0$ such that

$$E[\theta_\varepsilon^m(t) \theta^k(t)] \leq M_{m,k}, \forall t \in [0, T]. \tag{3.5}$$

We now can prove that $\theta_\varepsilon(t) \xrightarrow{L^2} \theta(t)$ uniformly with respect to $t \in [0, T]$, i.e:

$$\lim_{\varepsilon \rightarrow 0} \sup_{0 \leq t \leq T} \|\theta_\varepsilon(t) - \theta(t)\|_2 = 0. \tag{3.6}$$

Indeed, we see that

$$\begin{aligned} \|\theta_\varepsilon(t) - \theta(t)\|_2 &\leq \|\theta(t)\|_4 \|\exp(-\frac{1}{2}c^2\varepsilon^{2\alpha}t + c(\tilde{B}_t - B_t)) - 1\|_4 \\ &\leq M_4 \|\exp(-\frac{1}{2}c^2\varepsilon^{2\alpha}t + c(\tilde{B}_t - B_t)) - 1\|_4. \end{aligned} \tag{3.7}$$

Using the relation $e^x - 1 = x + o(x)$, we obtain

$$\begin{aligned} &\|\exp(-\frac{1}{2}c^2\varepsilon^{2\alpha}t + c(\tilde{B}_t - B_t)) - 1\|_4 \\ &\leq \|\frac{1}{2}c^2\varepsilon^{2\alpha}t + c(\tilde{B}_t - B_t)\|_4 + \|o(\dots)\|_4 \\ &\leq \frac{1}{2}c^2\varepsilon^{2\alpha}T + \|c(\tilde{B}_t - B_t)\|_4 + \|o(\dots)\|_4 \end{aligned} \tag{3.8}$$

and thus (3.6) follows from Corollary 2.2.

We have also that $\int_0^t \theta_\varepsilon^{n-1}(s)ds \xrightarrow{L^2} \int_0^t \theta^{n-1}(s)ds$ uniformly with respect to $t \in [0, T]$. Indeed, we have the following estimates:

$$\begin{aligned} &E \left| \int_0^t \theta_\varepsilon^{n-1}(s)ds - \int_0^t \theta^{n-1}(s)ds \right|^2 \\ &\leq t \int_0^t E|\theta_\varepsilon^{n-1}(t) - \theta^{n-1}(t)|^2 ds \forall t \in [0, T]. \end{aligned} \tag{3.9}$$

An application of Hölder inequality again yields for every $t \in [0, T]$

$$\begin{aligned} & E|\theta_\varepsilon^{n-1}(t) - \theta^{n-1}(t)|^2 \\ &= E[|\theta_\varepsilon(t) - \theta(t)| A_\varepsilon(t)] \\ &\leq \|\theta_\varepsilon(t) - \theta(t)\|_2 \|A_\varepsilon(t)\|_2, \end{aligned} \tag{3.10}$$

where $A_\varepsilon(t) = |\theta_\varepsilon(t) - \theta(t)|(\theta_\varepsilon^{n-2}(t) + \theta_\varepsilon^{n-3}(t)\theta(t) + \dots + \theta^{n-2}(t))^2$.

Using inequalities of the form (3.5) we see that there exists a finite constant $M_n > 0$ such that

$$\|A_\varepsilon(t)\|_2 \leq M_n, \quad \forall t \in [0, T]. \tag{3.11}$$

It follows from (3.9),(3.10) and (3.11) that

$$\begin{aligned} & E \left| \int_0^t \theta_\varepsilon^{n-1}(s) ds - \int_0^t \theta^{n-1}(s) ds \right|^2 \\ &\leq M_n t^2 \sup_{0 \leq t \leq T} \|\theta_\varepsilon(t) - \theta(t)\|_2 \\ &\leq M_n T^2 \sup_{0 \leq t \leq T} \|\theta_\varepsilon(t) - \theta(t)\|_2 \quad \forall t \in [0, T]. \end{aligned} \tag{3.12}$$

The last inequality assures that

$$\sup_{0 \leq t \leq T} \left\| \int_0^t \theta_\varepsilon^{n-1}(s) ds - \int_0^t \theta^{n-1}(s) ds \right\|_2 \rightarrow 0 \quad \text{as } \varepsilon \rightarrow 0.$$

Put $\eta_\varepsilon(t) = X_0^{1-n} + a(1-n) \int_0^t \theta_\varepsilon^{n-1}(s) ds$ and $\eta(t) = X_0^{1-n} + a(1-n) \int_0^t \theta^{n-1}(s) ds$.

From the results above we can see that $\eta_\varepsilon(t) \xrightarrow{L^2} \eta(t)$ uniformly with respect to $t \in [0, T]$. Next we will show that

$$\eta_\varepsilon^{\frac{1}{1-n}}(t) \xrightarrow{L^2} \eta^{\frac{1}{1-n}}(t) \quad \text{uniformly with respect to } t \in [0, T]. \tag{3.13}$$

Indeed, since $a \leq 0$, we have $\eta_\varepsilon(t) \geq X_0^{1-n}$ and $\eta(t) \geq X_0^{1-n}$ a.s. for every $t \in [0, T]$.

The theorem of finite increments applied to the function $g(x) = x^{\frac{1}{1-n}}$ yields

$$\left\| \eta_\varepsilon^{\frac{1}{1-n}}(t) - \eta^{\frac{1}{1-n}}(t) \right\|_2 \leq \frac{1}{n-1} \|X_0^n(\eta_\varepsilon(t) - \eta(t))\|_2.$$

By an argument analogous to the previous one, we get

$$\left\| \eta_\varepsilon^{\frac{1}{1-n}}(t) - \eta^{\frac{1}{1-n}}(t) \right\|_2 \leq M \|\eta_\varepsilon(t) - \eta(t)\|_2, \quad \forall t \in [0, T],$$

where $M > 0$ is a finite constant. And (3.13) now follows from this estimate.

As a consequence we have the following assertion

$$\tilde{X}_t = \theta_\varepsilon(t) \eta_\varepsilon^{\frac{1}{1-n}}(t) \xrightarrow{L^2} \theta(t) \eta^{\frac{1}{1-n}}(t) = X^*(t).$$

The proof of the theorem is thus complete. ■

4. Applications

We can check that some famous equations can be considered as particular cases of our fractional equation studied in this paper.

1. The fractional Black-Scholes model given by

$$dX_t = \mu X_t dt + \nu X_t dB_t,$$

$$a = 0, b = \mu, c = \nu,$$

$$X_t = X_0 e^{\nu B_t + \mu t}.$$

2. The fractional Verhulst equation

$$dX_t = (-X_t^2 + \lambda X_t) dt + \sigma X_t dB_t,$$

$$a = -1, b = \lambda, c = \sigma, n = 2,$$

$$X_t = e^{\sigma B_t + \lambda t} \left(X_0^{-1} + (-1)(-1) \int_0^t e^{\sigma B_s + \lambda s} ds \right)^{-1},$$

$$X_t = \frac{X_0 e^{\sigma B_t + \lambda t}}{1 + X_0 \int_0^t e^{\sigma B_s + \lambda s} ds}.$$

3. The fractional Ginzburg-Landau equation

$$dX_t = \left(-X_t^3 + \left(\alpha + \frac{\sigma^2}{2} \right) X_t \right) dt + \sigma X_t dB_t,$$

$$a = -1, b = \alpha + \frac{\sigma^2}{2}, c = \sigma, n = 3,$$

$$X_t = \frac{X_0 e^{\sigma B_t + (\alpha + \frac{\sigma^2}{2})t}}{\left(1 + 2X_0^2 \int_0^t e^{\sigma B_s + (\alpha + \frac{\sigma^2}{2})s} ds \right)^{\frac{1}{2}}}.$$

References

1. E. Alòs, O. Mazet, and D. Nualart, Stochastic calculus with respect to fractional Brownian motion with Hurst parameter less than $\frac{1}{2}$, *J. Stochastic processes and their applications* **86** (2000), 121–139.
2. F. Comte and E. Renault, Long Memory in Continuous-Time Stochastic Volatility Models, *Mathematical Finance* **8** (1998), 291–323.
3. K. Dębicki, Z. Michna, and T. Rolski, On the supremum from Gaussian processes over infinite horizon, *Probability and Math. Stat.* **18** (1998), 83–100.
4. L. Decreusefond and A. S. Üstünel, Stochastic Analysis of the Fractional Brownian Motion, *J. Potential Analysis* **10** (1999), 177–214.
5. T. E. Duncan, Y. Hu, and B. Pasik Duncan, Stochastic Calculus for Fractional Brownian Motion, *SIAM Control and Optimization* **38** (2000), 582–612.
6. D. Feyel de la A. Pradelle, Fractional integrals and Brownian processes, *Potential Analysis* **10** (1996) 273–288.
7. J. Jacques and R. Manca, *Semi-Markov Risk Models For Finance*, Insurance And Reliability, Springer, 2007.
8. B. Mandelbrot and van J. Ness, Fractional Brownian motions, Fractional Noises and Applications, *J. SIAM Review* **10** (1968), 422–437.
9. Tran Hung Thao, An approximate approach to fractional analysis for finance, *Nonlinear Analysis* **7** (2006), 124–132.